

Asymptotic Detection Performance of the Robust ANMF

Frédéric Pascal¹ and Jean-Philippe Ovarlez^{2,3}

¹L2S, CentraleSupélec, France

²SONDRA, CentraleSupélec, France

³French Aerospace Lab, ONERA DEMR/TSI, France

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Motivations

Let us considering the following binary hypotheses test:

$$\begin{cases} H_0 : \mathbf{y} = \mathbf{c}, & \mathbf{y}_i = \mathbf{c}_i, \quad i = 1, \dots, N \\ H_1 : \mathbf{y} = \alpha \mathbf{p} + \mathbf{c}, & \mathbf{y}_i = \mathbf{c}_i, \quad i = 1, \dots, N \end{cases},$$

where \mathbf{c} is an additive noise, $\{\mathbf{c}_i\}_{i \in [1, N]}$ are N signal-free secondary data, \mathbf{p} a known steering vector and where α is the unknown amplitude of the target.

In partially homogeneous Gaussian environment (i.e. $\{\mathbf{c}_i\}_{i \in [1, N]} \sim \mathbb{CN}(\mathbf{0}_m, \mathbf{M})$, $\mathbf{c} \sim \mathbb{CN}(\mathbf{0}_m, \sigma^2 \mathbf{M})$) when \mathbf{M} is known and σ^2 unknown, the GLRT is the well known **Normalized Matched Filter** [L. Scharf]:

$$H(\mathbf{M}) = \frac{|\mathbf{p}^H \mathbf{M}^{-1} \mathbf{y}|^2}{(\mathbf{p}^H \mathbf{M}^{-1} \mathbf{p})(\mathbf{y}^H \mathbf{M}^{-1} \mathbf{y})}.$$

Motivations

When an estimate $\widehat{\mathbf{M}}_{SCM} = \frac{1}{N} \sum_{k=1}^N \mathbf{c}_k \mathbf{c}_k^H$ of \mathbf{M} is plugged into the NMF (two-step GLRT), this results in the so-called ANMF [S. Kraut] whose distribution is given by:

$$p_{H(\widehat{\mathbf{M}}_{SCM})}(x) = \frac{\Gamma(N+1) e^{-\delta}}{\Gamma(N-m+1) \Gamma(m-1)} \int_0^1 u^{N-m+1} \frac{(1-u)^{m-1} (1-x)^{N-m}}{(1-ux)^{N-m+2}} \times {}_1F_1 \left(N-m+2, 1; \frac{\delta x (1-u)}{1-xu} \right) du. \quad (1)$$

Performances of the NMF and ANMF can be easily described in terms of:

- Probability of False Alarm P_{fa} versus the detection threshold λ ($\delta = 0$),
- Probability of Detection P_d versus the SNR $\delta = \alpha^2 \mathbf{p}^H \mathbf{M}^{-1} \mathbf{p} / \sigma^2$.

Motivations

Under more severe environment (spikyness, heterogeneity of the background, outliers in secondary data, ...), the performance of ANMF is dramatically degraded. The noise \mathbf{c} and secondary data $\{\mathbf{c}_i\}_{i \in [1, N]}$ cannot be described by conventional Gaussian PDF:

- need to characterize the environment statistics using more general models: Spherically Invariant Random Vectors (SIRV) or Complex Elliptically Symmetric (CES) distributions
- need to propose robust estimators of the background parameters (e.g. covariance matrix): M -estimators

The goal of this paper is to derive under both H_0 and H_1 hypotheses the asymptotic distributions (N not too small) of the robust ANMF built with any M -estimator $\widehat{\mathbf{M}}$ when the noise and secondary data are modelled by Complex Elliptically Symmetric (CES) distributions:

$$H(\widehat{\mathbf{M}}) = \frac{\left| \mathbf{p}^H \widehat{\mathbf{M}}^{-1} \mathbf{y} \right|^2}{\left(\mathbf{p}^H \widehat{\mathbf{M}}^{-1} \mathbf{p}^H \right) \left(\mathbf{y}^H \widehat{\mathbf{M}}^{-1} \mathbf{y} \right)}.$$

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Modeling the Background

Let \mathbf{c} be a complex circular random vector of length m . \mathbf{c} has a complex elliptically symmetric (CES) distribution ($CE(\boldsymbol{\mu}, \mathbf{M}, h_m)$) if its PDF is

$$g_{\mathbf{c}}(\mathbf{c}) = |\mathbf{M}|^{-1} h_m \left((\mathbf{c} - \boldsymbol{\mu})^H \mathbf{M}^{-1} (\mathbf{c} - \boldsymbol{\mu}) \right),$$

where $h_m : [0, \infty) \rightarrow [0, \infty)$ is the density generator.

- $\boldsymbol{\mu}$ is the statistical mean (generally known or $= \mathbf{0}$)
- \mathbf{M} the scatter matrix

In general (finite second-order moment), \mathbf{M} is equal to the covariance matrix $E[(\mathbf{c} - \boldsymbol{\mu})(\mathbf{c} - \boldsymbol{\mu})^H]$ up to a scalar factor.

Attractive clutter modeling

Some important properties

- Large class of distributions: Gaussian, SIRV, MGGD, K-dist., Student-t....
- Closed under affine transformations,
- All sub-vectors of \mathbf{z} have a CES distribution,
- CES stochastic representation theorem: $\mathbf{c} =_d \boldsymbol{\mu} + \tau \mathbf{A} \mathbf{u}$ where the random scalar texture $\tau \geq 0$ is independent of \mathbf{u} (m -vector uniformly distributed on the sphere) and characterized by its PDF $p_\tau(\cdot)$ and where $\mathbf{M} = \mathbf{A} \mathbf{A}^H$,
- SIRV subclass stochastic representation theorem: $\mathbf{c} =_d \boldsymbol{\mu} + \tau \mathbf{u}$ where the random scalar texture $\tau \geq 0$ is independent of $\mathbf{u} \sim \mathcal{CN}(\mathbf{0}_m, \mathbf{M})$ and characterized by its PDF $p_\tau(\cdot)$.

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Estimating the covariance matrix

Let $(\mathbf{c}_1, \dots, \mathbf{c}_N)$ be a N -sample $\sim \text{CES}(\mathbf{0}_m, \mathbf{M}, h_m)$ (Secondary data).

M -estimator of \mathbf{M}

$$\widehat{\mathbf{M}} = \frac{1}{N} \sum_{i=1}^N u \left(\mathbf{c}_i^H \widehat{\mathbf{M}}^{-1} \mathbf{c}_i \right) \mathbf{c}_i \mathbf{c}_i^H,$$

Maronna (1976), Kent and Tyler (1991)

- Existence
- Uniqueness
- Convergence of the recursive algorithm...

PDF specified \Rightarrow MLE can be derived: $u(x) = -h'_m(x)/h_m(x)$

PDF not specified \Rightarrow general M -estimators are used instead

Some remarks on the \mathbf{M} -estimators

- FPE and SCM are “not” (theoretically) \mathbf{M} -estimators
- FPE is the most robust

FP Estimate (Tyler, 1987; Pascal, 2008)

$$\widehat{\mathbf{M}}_{FPE} = \frac{m}{N} \sum_{i=1}^N \frac{\mathbf{c}_i \mathbf{c}_i^H}{\mathbf{c}_i^H \widehat{\mathbf{M}}_{FPE}^{-1} \mathbf{c}_i}$$

- The FPE does not depend on the texture (SIRV or CES distributions)
- Existence, Uniqueness, Convergence of the recursive algorithm...
- True MLE under SIRV noise with unknown deterministic texture $\{\boldsymbol{\tau}_i\}_{i \in [1, N]}$

Asymptotic distribution of complex M -estimators

Using the results of Tyler, we derived the following results [Ollila, Mahot, 2013]:

Asymptotic distribution of $\widehat{\mathbf{M}}$

$$\sqrt{N} \operatorname{vec}(\widehat{\mathbf{M}} - \mathbf{M}) \xrightarrow{d} \mathbb{CN}(\mathbf{0}_{m^2}, \mathbf{C}, \mathbf{P}),$$

where the covariance matrix \mathbf{C} and the pseudo covariance matrix \mathbf{P} are given by:

$$\begin{aligned} \mathbf{C} &= \mathbf{v}_1(\mathbf{M}^* \otimes \mathbf{M}) + \mathbf{v}_2 \operatorname{vec}(\mathbf{M}) \operatorname{vec}(\mathbf{M})^H, \\ \mathbf{P} &= \mathbf{v}_1(\mathbf{M}^* \otimes \mathbf{M}) \mathbf{K} + \mathbf{v}_2 \operatorname{vec}(\mathbf{M}) \operatorname{vec}(\mathbf{M})^T, \end{aligned}$$

where \mathbf{K} is the commutation matrix and where the constant \mathbf{v}_1 and \mathbf{v}_2 are completely defined.

Asymptotic distribution of a function of complex \mathbf{M} -estimators

- Let $H(\mathbf{V})$ be a r -multivariate function on the set of Hermitian positive-definite matrices, with continuous first partial derivatives and such as $H(\mathbf{V}) = H(\alpha \mathbf{V})$ for all $\alpha > 0$, e.g. the ANMF statistic, the MUSIC statistic.

Asymptotic distribution of $H(\mathbf{M})$

$$\sqrt{N} \left(H(\widehat{\mathbf{M}}) - H(\mathbf{M}) \right) \xrightarrow{d} \mathcal{CN}(\mathbf{0}_r, \mathbf{C}_H, \mathbf{P}_H)$$

where \mathbf{C}_H and \mathbf{P}_H are defined as

$$\begin{aligned} \mathbf{C}_H &= \mathbf{v}_1 H'(\mathbf{M}) (\mathbf{M}^T \otimes \mathbf{M}) H'(\mathbf{M})^H, \\ \mathbf{P}_H &= \mathbf{v}_1 H'(\mathbf{M}) (\mathbf{M}^T \otimes \mathbf{M}) \mathbf{K}_{m,m} H'(\mathbf{M})^T, \end{aligned}$$

where $H'(\mathbf{M}) = \left(\frac{\partial H(\mathbf{M})}{\partial \text{vec}(\mathbf{M})} \right)$.

Two important properties of complex M -estimators

- SCM and M -estimators share the same asymptotic distribution
- $H(SCM)$ and $H(M\text{-estimators})$ share the same asymptotic distribution (differs from ν_1)

	SCM	M -estimators	FP
ν_1	1	ν_1	$(m+1)/m$
ν_2	0	ν_2	$-(m+1)/m^2$

This important result shows that asymptotically and under Gaussian environment:

- any M -estimator built with N observations behaves like the SCM but with a slight smaller degree of freedom N/ν_1 ,
- any function H built with M -estimator behaves like those built with SCM but with a slight smaller degree of freedom N/ν_1 .

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A particular choice of $H(\cdot)$: the two-step GLRT ANMF

Robust ANMF test (ACE, GLRT-LQ) [Conte, 1995]

$$H(\widehat{\mathbf{M}}) = \Lambda_{ANMF}(\mathbf{y}, \widehat{\mathbf{M}}) = \frac{|\mathbf{p}^H \widehat{\mathbf{M}}^{-1} \mathbf{y}|^2}{(\mathbf{p}^H \widehat{\mathbf{M}}^{-1} \mathbf{p})(\mathbf{y}^H \widehat{\mathbf{M}}^{-1} \mathbf{y})} \underset{H_0}{\overset{H_1}{\gtrless}} \lambda$$

where $\widehat{\mathbf{M}}$ stands for any M -estimators.

- The ANMF is **scale-invariant** (homogeneous of degree 0), i.e.

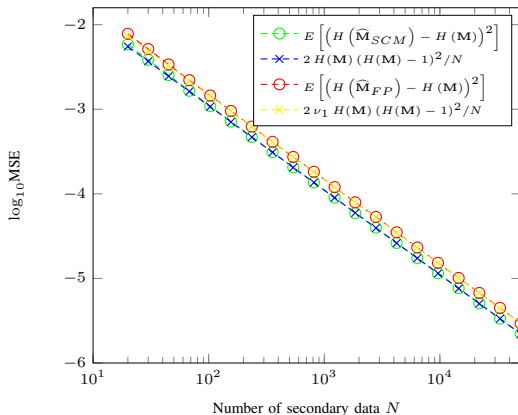
$$\forall \alpha, \beta \in \mathbb{R}, \Lambda_{ANMF}(\alpha \mathbf{y}, \beta \widehat{\mathbf{M}}) = \Lambda_{ANMF}(\mathbf{y}, \widehat{\mathbf{M}})$$

- The ANMF test is CFAR w.r.t the covariance/scatter matrix \mathbf{M} ,
- The ANMF test is CFAR w.r.t the texture (SIRV or CES distributions)

Two different ways to derive robust ANMF asymptotic performance

- by **correcting the degree of freedom** of the M -estimator ($N \rightarrow N/\nu_1$) in $p_{H(\widehat{\mathbf{M}}_{SCM})}$ presented in (1) and by conditioning on the texture PDF,
- by exploiting directly the **asymptotic distribution** of the ANMF [Pascal -Ovarlez 2015]:

$$H(\widehat{\mathbf{M}}) \xrightarrow{d} \mathbb{CN} \left(H(\mathbf{M}), 2 \frac{\nu_1}{N} H(\mathbf{M}) (H(\mathbf{M}) - 1)^2 \right).$$



Empirical variance of the ANMF built with the SCM ($\nu_1 = 1$) and Tyler's M -estimator ($\nu_1 = (m + 1)/m$) in Gaussian environment and theoretical asymptotic variance for $m = 3$ and $\mathbf{M} = \mathbf{I}_3$.

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Correcting the degree of freedom of the M -estimator

Evaluation of the performance for a cell under test containing SIRV noise and for any CES secondary data.

$$P_{fa} = \mathbb{P} \left(H(\widehat{\mathbf{M}}) \geq \lambda | H_0 \right), \quad P_d = \mathbb{P} \left(H(\widehat{\mathbf{M}}) \geq \lambda | H_1 \right)$$

$$P_{fa} = (1 - \lambda)^{N/\nu_1 - m + 1} {}_2F_1 \left(N/\nu_1 - m + 2, N/\nu_1 - m + 1; N/\nu_1 + 1; \lambda \right),$$

$$P_d = 1 - \int_0^{+\infty} d\tau \int_0^1 du \int_0^\lambda u^{N/\nu_1 - m + 1} \frac{(1 - u)^{m-1} (1 - x)^{N/\nu_1 - m}}{(1 - ux)^{N/\nu_1 - m + 2}} e^{-\delta/\tau} \\ \times \frac{\Gamma(N/\nu_1 + 1)}{\Gamma(N/\nu_1 - m + 1) \Gamma(m - 1)} {}_1F_1 \left(N/\nu_1 - m + 2, 1; \frac{\delta}{\tau} \frac{x(1 - u)}{1 - xu} \right) p_\tau(\tau) dx$$

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Exploiting the asymptotic distribution of the robust ANMF

$$H(\widehat{\mathbf{M}}) \xrightarrow{d} \mathcal{CN} \left(H(\mathbf{M}), 2 \frac{\nu_1}{N} H(\mathbf{M}) (H(\mathbf{M}) - 1)^2 \right)$$

The distribution of $H(\widehat{\mathbf{M}})$ is conditioned to the distribution $p_{H(\mathbf{M})}$ of $H(\mathbf{M})$:

- the cell under test contains Gaussian noise:

$$p_{H(\mathbf{M})}(u) = e^{-\delta} \beta_{1,m-1}(u) {}_1F_1(m, 1; u\delta).$$

- the cell under test contains SIRV noise:

$$p_{H(\mathbf{M})}(u) = \int_0^\infty e^{-\delta/\tau} \beta_{1,m-1}(u) {}_1F_1 \left(m, 1; \frac{u\delta}{\tau} \right) p_\tau(\tau) d\tau.$$

- the cell under test contains general CES noise: No closed-form under H_1 .
 The same as SIRV noise under H_0 .

Exploiting the asymptotic distribution of the robust ANMF

Evaluation of the performance for **a cell under test containing Gaussian noise** and for **any CES secondary data**.

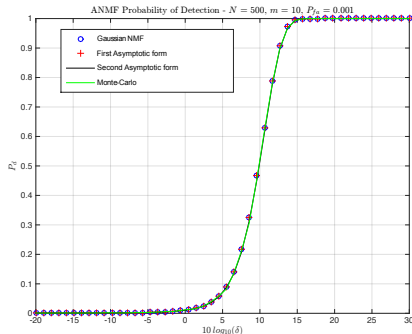
$$P_{fa} = \mathbb{P} \left(H(\widehat{\mathbf{M}}) \geq \lambda | H_0 \right), \quad P_d = \mathbb{P} \left(H(\widehat{\mathbf{M}}) \geq \lambda | H_1 \right)$$

$$P_{fa} = 1 - \int_0^1 \beta_{1,m-1}(x) \Phi \left(\frac{\sqrt{N}(\lambda - x)}{\sqrt{2\nu_1 x(x-1)^2}} \right) dx.$$

$$P_d = 1 - \int_0^1 \beta_{1,m-1}(x) e^{\delta(x-1)} {}_1F_1(1-m, 1; -x\delta) \Phi \left(\frac{\sqrt{N}(\lambda - x)}{\sqrt{2\nu_1 x(x-1)^2}} \right) dx$$

where $\Phi(\cdot)$ is the cumulative distribution of the Normal distribution.

Ex: K-distributed secondary data and Gaussian noise in the cell under test.



Comparison between P_d and SNR δ relationships for the ANMF built with Tyler's estimator, $m = 10$, $N = 500$ and $P_{fa} = 10^{-3}$, $\mathbf{p} = [1, \dots, 1]^T$, $\{\mathbf{y}_i\}_{i \in [1, N]} \sim K_v$ where K_v is a K-distribution with shape $v = 0.1$. and $\mathbf{y} \sim \mathcal{CN}(\alpha \mathbf{p}, \mathbf{M})$ where $\mathbf{M} = (\rho^{|i-j|})_{i,j}$ with $\rho = 0.5$

Exploiting the asymptotic distribution of the robust ANMF

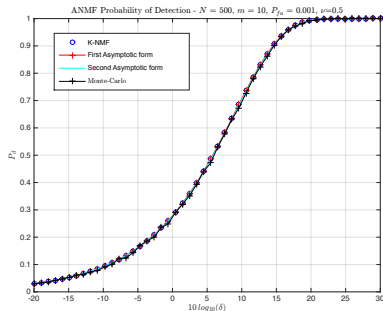
Evaluation of the performance for a cell under test containing SIRV noise with texture $\sim p_\tau(\cdot)$ and for any CES secondary data.

$$P_{fa} = \mathbb{P} \left(H(\widehat{\mathbf{M}}) \geq \lambda | H_0 \right), \quad P_d = \mathbb{P} \left(H(\widehat{\mathbf{M}}) \geq \lambda | H_1 \right)$$

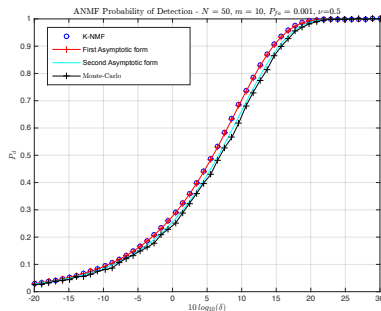
$$P_{fa} = 1 - \int_0^1 \beta_{1,m-1}(x) \Phi \left(\frac{\sqrt{N}(\lambda - x)}{\sqrt{2\nu_1 x(x-1)^2}} \right) dx.$$

$$P_d = 1 - \int_0^\infty p_\tau(\tau) d\tau \int_0^1 \beta_{1,m-1}(x) e^{\delta(x-1)/\tau} {}_1F_1 \left(1-m, 1; -x \frac{\delta}{\tau} \right) \\ \times \Phi \left(\frac{\sqrt{N}(\lambda - x)}{\sqrt{2\nu_1 x(x-1)^2}} \right) dx$$

where $\Phi(\cdot)$ is the cumulative distribution of the Normal distribution.



$N = 500$



$N = 50$

Comparison between P_d and SNR δ relationships for the NMF, the ANMF built with Tyler's estimator and its asymptotic form, $m = 10$, $\nu_1 = 1.1$ and $P_{fa} = 10^{-3}$, $\mathbf{p} = [1, \dots, 1]^T$, $\mathbf{y} = \alpha \mathbf{p} + \mathbf{c}$ where $\mathbf{c} \sim K_\nu$ where K_ν is a multivariate K-distribution with shape parameter $\nu = 0.5$ and covariance matrix $\mathbf{M} = \left(\rho^{|i-j|} \right)_{i,j}$ with $\rho = 0.5$.

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Conclusions

Two asymptotic approximations of the corresponding robust ANMF distribution have been derived following different approaches:

- First asymptotic distribution is based on the correction of the degrees of freedom of M -estimators
- Second asymptotic distribution is based on the direct exploitation of the asymptotic distribution of the ANMF

These results provide a very good approximation of the robust ANMF distribution in CES environment even for a small number of observations and have been applied to theoretically regulate the false alarm probability and to evaluate the detection performance.