

# Robust Statistics for Multi-Band SAR Image Change Detection

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# Plan

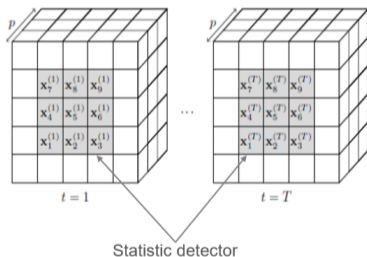
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- Introduction
  - Background on change detection methods in SAR imagery
  - Problems in multivariate and multiband SAR image analysis.
  - Contributions in multiband SAR image change detection
- Proposed multiband change detection method
  - Multiband SAR modeling
  - Derivation of the detector
  - Applications
    - Simulated data - False Alarm Regulation
    - Simulated data - Probability of Detection (PD) vs Signal-to-Noise -Ratio (SNR)
    - Real data - L and X bands SETHI polarimetric SAR images.
- Conclusions and future works

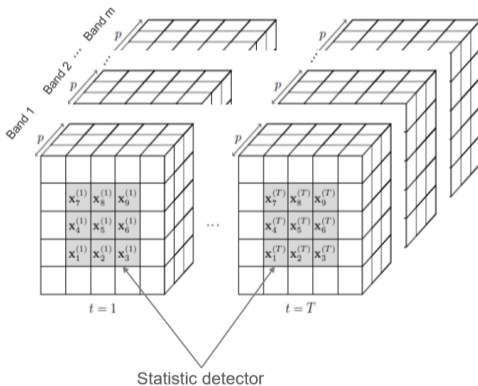
# Introduction

## Background- Detection concept for polarimetric SAR image [1]

- Mono-band



- Multiband



In our case  $T = 2$ ,  $x_5^{(t)}$  is the test pixel at time  $t$ .

# Introduction

## Background - Conventional Conradsen's detector [2] on single-band

The vector  $\mathbf{x} \in \mathbb{C}^p$  is distributed according to complex circular Normal  $\mathcal{CN}(\mathbf{0}_p, \Phi)$  distribution to handle spatial SAR image homogeneity:

$$p_{\mathbf{x}}(\mathbf{x}, \Phi) = \frac{1}{\pi^p |\Phi|} \exp(-\mathbf{x}^H \Phi^{-1} \mathbf{x}),$$

where  $\Phi$  is the unknown covariance matrix.

$$\text{Detection problem: } \begin{cases} H_0 : \Phi_1 = \Phi_2 \\ H_1 : \Phi_1 \neq \Phi_2 \end{cases}$$

The Conradsen's classical Gaussian detector [2] is given by:

$$\hat{\Lambda}_G = \frac{|\hat{\Phi}^{SCM}|^{2N}}{\prod_{t=1}^{t=2} |\hat{\Phi}_t^{SCM}|^N} \underset{H_0}{\overset{H_1}{\gtrless}} \lambda, \quad \text{where } \hat{\Phi}^{SCM} \text{ and } \hat{\Phi}_t^{SCM} \text{ are Sample Covariance Matrices.}$$



# Introduction

## Background - A. Mian Detector [1] on single-band

The vector  $\mathbf{x} \in \mathbb{C}^p$  is distributed according to  $\mathcal{CN}(\mathbf{0}_p, \tau \Phi)$  to handle spatial SAR image heterogeneity.

$$p_{\mathbf{x}}(\mathbf{x}, \Phi) = \frac{1}{\pi^p \tau^p |\Phi|} \exp\left(-\frac{\mathbf{x}^H \Phi^{-1} \mathbf{x}}{\tau}\right),$$

where the covariance matrix  $\Phi$  and the positive scalar texture  $\tau$  are unknown parameters.

$$\text{Detection problem: } \begin{cases} H_0 : \{\Phi_1, \tau_1\} = \{\Phi_2, \tau_2\} \\ H_1 : \{\Phi_1, \tau_1\} \neq \{\Phi_2, \tau_2\} \end{cases}$$

The Mian's GLRT Detector [1] is defined as:

$$\hat{\Lambda}_{MT} = \frac{|\hat{\Phi}^{FP}|^{2N}}{\prod_{t=1}^N |\hat{\Phi}_t^{FP}|^N} \prod_{k=1}^N \frac{|\hat{\tau}_k|^4}{\prod_{t=1}^N |\hat{\tau}_k^t|^2} \underset{H_0}{\overset{H_1}{\gtrless}} \lambda, \quad \text{where } \hat{\Phi}^{FP} \text{ and } \hat{\Phi}_t^{FP} \text{ are Tyler Covariance Matrices.}$$

# Introduction

## Problems and Contributions

### Identified problems for multi-band case:

- Textures (or heterogeneity) are not taken into account in the design of the Gaussian detector  $\hat{\Lambda}_G$ .
- The detector  $\hat{\Lambda}_{MT}$  will perform better only if the concatenated bands are characterized by the same texture.

### Contributions:

- Proposal for a multi-band change detection detector that considers heterogeneous nature and variations in texture across the merged bands (inspired by [3]).
- Applications on simulated data to analyse regulation of False Alarm and performance in terms of PD/SNR.
- Applications on polarimetric SAR data simultaneously acquired at two bands, L and X, by the ONERA SETHI SAR system at two different dates,  $t_1$  and  $t_2$  at Captieux.

# Proposed multiband change detection method

## Multiband SAR modeling

The vector  $\mathbf{x} = [\mathbf{x}_1^T, \mathbf{x}_2^T, \dots, \mathbf{x}_m^T]^T \in \mathbb{C}^{mp}$  is distributed as  $\mathcal{CN}(\mathbf{0}_{mp}, \mathbf{\Sigma})$ :

$$p_{\mathbf{x}}(\mathbf{x}) = \frac{1}{\pi^{mp} |\mathbf{\Sigma}|} \exp(-\mathbf{x}^H \mathbf{\Sigma}^{-1} \mathbf{x}),$$

where  $\mathbf{x}_i = \sqrt{\tau_i} \mathbf{z}_i$  and with  $\mathbf{\Sigma} = E(\mathbf{x} \mathbf{x}^H) = \mathbf{T} \mathbf{\Phi} \mathbf{T}$  where

$$\mathbf{\Phi} = \begin{pmatrix} E(\mathbf{z}_1 \mathbf{z}_1^H) & E(\mathbf{z}_1 \mathbf{z}_2^H) & \dots & E(\mathbf{z}_1 \mathbf{z}_m^H) \\ E(\mathbf{z}_2 \mathbf{z}_1^H) & E(\mathbf{z}_2 \mathbf{z}_2^H) & \dots & E(\mathbf{z}_2 \mathbf{z}_m^H) \\ \vdots & \vdots & \ddots & \vdots \\ E(\mathbf{z}_m \mathbf{z}_1^H) & E(\mathbf{z}_m \mathbf{z}_2^H) & \dots & E(\mathbf{z}_m \mathbf{z}_m^H) \end{pmatrix}, \mathbf{T} = \begin{pmatrix} \sqrt{\tau_1} \mathbf{I}_p & \mathbf{0}_p & \dots & \mathbf{0}_p \\ \mathbf{0}_p & \sqrt{\tau_2} \mathbf{I}_p & \dots & \mathbf{0}_p \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{0}_p & \mathbf{0}_p & \dots & \sqrt{\tau_m} \mathbf{I}_p \end{pmatrix}.$$

The change detection problem is then characterized by:  $\begin{cases} H_0 : \mathbf{\Sigma}_1 = \mathbf{\Sigma}_2 \\ H_1 : \mathbf{\Sigma}_1 \neq \mathbf{\Sigma}_2 \end{cases}$ .

# Proposed multiband change detection method

## Derivation of the detector - Estimation by the joint fixed-point method

The parameters  $\left( \hat{\Phi}, \left\{ \hat{\mathbf{T}}_k \right\}_{k \in \{1, N\}} \right)$  et  $\left( \left\{ \hat{\mathbf{T}}_k^t \right\}_{k \in \{1, N\}, t \in \{1, 2\}}, \left\{ \hat{\Phi}^t \right\}_{t \in \{1, 2\}} \right)$  are given:

- by the two following joint fixed point equations under  $H_0$ ,

$$\begin{cases} \hat{\mathbf{T}}_k = \frac{1}{2} \sum_{t=1}^2 \operatorname{Re} \left( \hat{\Phi}^{-1} \hat{\mathbf{T}}_k^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \right), \quad \forall k \in \{1, N\}, \\ \hat{\Phi} = \frac{1}{2N} \sum_{k=1, t=1}^{k=N, t=2} \hat{\mathbf{T}}_k^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \hat{\mathbf{T}}_k^{-1}. \end{cases}$$

- and by the two following joint fixed point equations under  $H_1$ ,

$$\begin{cases} \hat{\mathbf{T}}_k^t = \operatorname{Re} \left( \hat{\Phi}_t^{-1} \left( \hat{\mathbf{T}}_k^t \right)^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \right), \quad \forall k \in \{1, N\}, \forall t \in \{1, 2\} \\ \hat{\Phi}_t = \frac{1}{N} \sum_{k=1}^N \left( \hat{\mathbf{T}}_k^t \right)^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \left( \hat{\mathbf{T}}_k^t \right)^{-1}, \quad \forall t \in \{1, 2\} \end{cases}$$

# Proposed multiband change detection method

## Derivation of the detector

Estimating the parameters  $(\mathbf{T}_k, \Phi)$  under  $H_0$  and  $(\mathbf{T}_k^t, \Phi_t)$  under  $H_1$  with  $N$  secondary data through Generalized Maximum Likelihood Estimation procedure leads to:

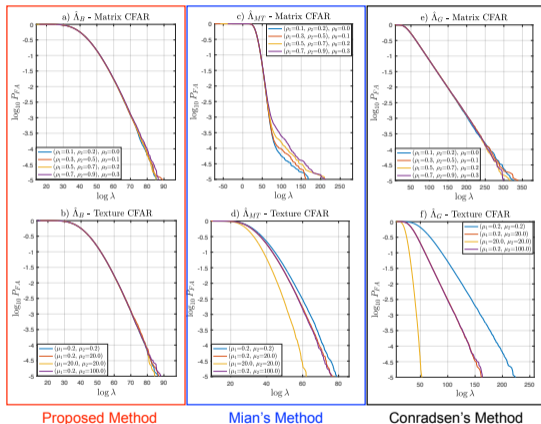
$$\hat{\Lambda}_B = \frac{|\hat{\Phi}|^{2N}}{\prod_{t=1}^N |\hat{\Phi}_t|^N} \prod_{k=1}^N \frac{|\hat{\mathbf{T}}_k|^4}{\prod_{t=1}^N |\hat{\mathbf{T}}_k^t|^2} \underset{H_0}{\overset{H_1}{\gtrless}} \lambda,$$

where  $\lambda$  is the detection threshold.

**Recall:** *CFAR (Constant False Alarm Regulation) Matrix or Texture property* means independence of the detection test with the PDF parameters!

# Results

## Simulated data - False Alarm Regulation



$P_{FA}$  vs  $\lambda$  with  $N = 16, p = 3, m = 2, k = 1$ .

- $\mathbf{x} = (\mathbf{x}_1^T, \mathbf{x}_2^T)^T, \mathcal{T}(\rho) = (\rho^{|i-j|})_{i,j}$

- Band 1:  $\mathbf{x}_1 = \sqrt{\tau_1} \mathbf{z}_1$

where  $\begin{cases} \mathbf{z}_1 \sim \mathcal{CN}(\mathbf{0}_p, \mathcal{T}(\rho_1)) \\ \tau_1 \sim \Gamma(\mu_1, 1/\mu_1) \end{cases}$ .

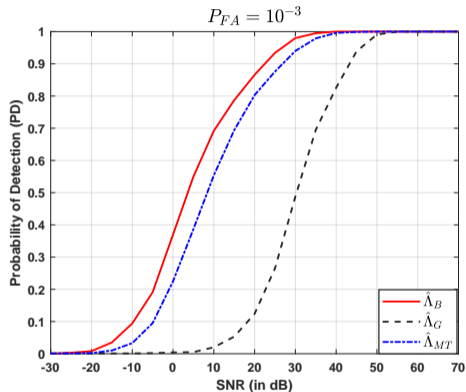
- Band 2:  $\mathbf{x}_2 = \sqrt{\tau_2} \mathbf{z}_2$

where  $\begin{cases} \mathbf{z}_2 \sim \mathcal{CN}(\mathbf{0}_p, \mathcal{T}(\rho_2)) \\ \tau_2 \sim \Gamma(\mu_2, k/\mu_2) \\ k \in \mathbb{R}^+ \text{ scale factor.} \end{cases}$ .

- Interband:  $E(\mathbf{z}_1 \mathbf{z}_2^H) = \rho_0 \mathbf{1}_p$ .

# Results

## Simulated data - Probability of Detection vs SNR (in dB)



Probability of detection vs SNR with textured Gaussian simulated data over two frequency bands ( $\rho_1 = 0.1$ ,  $\rho_2 = 0.5$ ,  $\rho_0 = 0.1$ ,  $\mu_1 = 1$  and  $\mu_2 = 0.1$ , with scale factor  $k = 50$  for  $P_{FA} = 10^{-3}$ ).

At PD = 0.7,  $\hat{\Lambda}_B$  gains approximately 5 and 25 dB compared to  $\hat{\Lambda}_{MT}$  and  $\hat{\Lambda}_G$ , respectively.

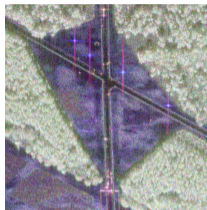
# Real data

## L and X band SETHI POLSAR images in Pauli basis in RGB color composition

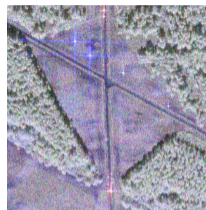


Ground truth of the studied scene

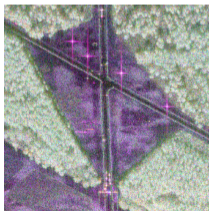
L-band -  $t_1$



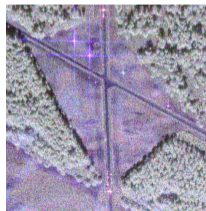
X-band -  $t_1$



L-band -  $t_2$

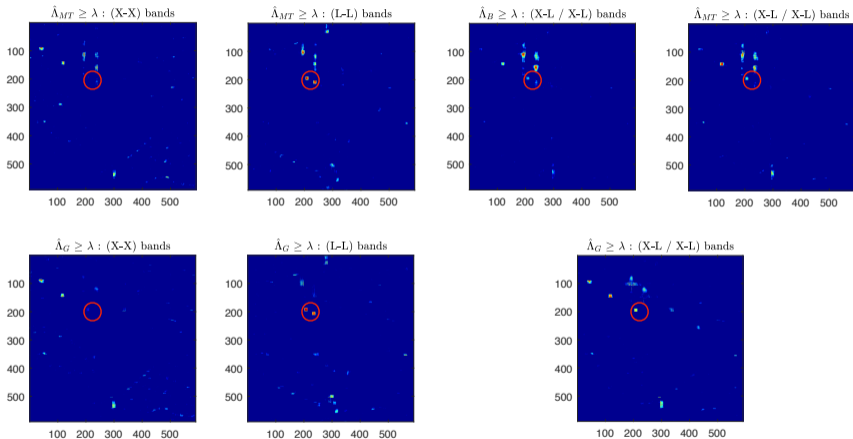


X-band -  $t_2$





# Results on experimental SETHI data



**Left:** Mono-band detectors  $\hat{\Lambda}_{MT}$  and  $\hat{\Lambda}_G$ . **Right:** L-X multi-band detectors  $\hat{\Lambda}_{MT}$ ,  $\hat{\Lambda}_G$  and  $\hat{\Lambda}_B$  ( $P_{fa} = 10^{-3}$ ).

# Conclusions and future works

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In this work, we have,

- proposed a robust statistic for multi-band SAR image change detection and analyzed his performance,
- applied the proposed detector on simulated data and SETHI SAR images in L and X bands at two dates  $t_1$  and  $t_2$ ,
  - the proposed detector are CFAR with respect to matrix and texture,
  - the proposed detector outperforms the state-of-the-art methods in terms of PD vs SNR and in terms of PFA regulation.
  - it shows promising results on real data when combining X-band and L-band.

In the future works, we will,

- analyze the convergence of the joint fixed point equations,
- analyze the performance on experimental data of the proposed detector on more than two bands.

# References

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- [1] A. Mian, G. Ginolhac, J.-P. Ovarlez, and A. M. Atto, “New robust statistics for change detection in time series of multivariate SAR images,” *IEEE Transactions on Signal Processing*, vol. 67, no. 2, pp. 520–534, 2019.
- [2] K. Conradsen, A. A. Nielsen, J. Schou, and H. Skriver, “Change detection in polarimetric SAR data and the complex Wishart distribution,” in *IGARSS 2001. Scanning the Present and Resolving the Future. Proceedings. IEEE 2001 International Geoscience and Remote Sensing Symposium (Cat. No.01CH37217)*, vol. 6, pp. 2628–2630 vol.6, 2001.
- [3] O. Lerda, A. Mian, G. Ginolhac, J.-P. Ovarlez, and D. Charlot, “Robust Detection for Mills Cross Sonar,” *IEEE Journal of Oceanic Engineering*, vol. 49, no. 3, pp. 1009–1024, 2024.

**Thank you for your attention!**  
**Any questions?**

# Appendix

Maximum Likelihood Estimation procedure with  $N$  secondary data to derive the detector:

$$\Lambda_B(\mathbf{x}) = \frac{\mathcal{L}_1 \left( \left\{ \{ \mathbf{x}_k^t, \mathbf{T}_k^t \}_{k \in [1, N]}, \Phi_t \right\}_{t \in [1, 2]} \right)}{\mathcal{L}_0 \left( \left\{ \{ \mathbf{x}_k^t \}_{k \in [1, N]} \right\}_{t \in [1, 2]}, \{ \mathbf{T}_k \}_{k \in [1, N]}, \Phi \right)},$$

$$\mathcal{L}_1 \left( \left\{ \{ \mathbf{x}_k^t, \mathbf{T}_k^t \}_k, \Phi_t \right\}_t \right) = \prod_{\substack{k=1 \\ t=1}}^{\substack{k=N \\ t=2}} p_{\mathbf{x}} \left( \mathbf{x}_k^t, \mathbf{T}_k^t \Phi_t \mathbf{T}_k^t \right),$$

$$\mathcal{L}_0 \left( \left\{ \{ \mathbf{x}_k^t \}_k \right\}_t, \{ \mathbf{T}_k \}_k, \Phi \right) = \prod_{\substack{k=1 \\ t=1}}^{\substack{k=N \\ t=2}} p_{\mathbf{x}} \left( \mathbf{x}_k^t, \mathbf{T}_k \Phi \mathbf{T}_k \right).$$

To maximize the two likelihood functions,  $\mathcal{L}_0$  and  $\mathcal{L}_1$ , we need to perform the following

# Appendix

To maximize the two likelihood functions,  $\mathcal{L}_0$  and  $\mathcal{L}_1$ , we need to perform the following operations: By denoting  $C = -2 m p N \log(\pi)$ , the logarithm of (15), under  $H_0$ , is given by:

$$\log(\mathcal{L}_0) = C - \sum_{\substack{k=1 \\ t=1}}^{\substack{k=N \\ t=2}} \log(|\Phi|) + 2 \sum_{\substack{k=1 \\ t=1}}^{\substack{k=N \\ t=2}} \log(|\mathbf{T}_k^{-1}|) - \sum_{\substack{k=1 \\ t=1}}^{\substack{k=N \\ t=2}} (\mathbf{x}_k^{tH} \mathbf{T}_k^{-1} \Phi^{-1} \mathbf{T}_k^{-1} \mathbf{x}_k^t) .$$

Now, we proceed by taking the derivative with respect to  $\mathbf{T}_k^{-1}$ . Then,  $\forall k \in [1, N]$ ,

$$\frac{\partial \log(\mathcal{L}_0)}{\partial \mathbf{T}_k^{-1}} = 4 \sum_{k=1}^{k=N} \mathbf{T}_k - 2 \sum_{\substack{k=1 \\ t=1}}^{\substack{k=N \\ t=2}} \text{Re}((\Phi \mathbf{T}_k)^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH}) = \mathbf{0}_{mp} .$$

# Appendix

As  $\frac{\partial \log (|\mathbf{T}_k^{-1}|)}{\partial \mathbf{T}_k^{-1}} = \mathbf{T}_k$  and  $\mathbf{x}_k^{tH} (\mathbf{T}_k \boldsymbol{\Phi} \mathbf{T}_k)^{-1} \mathbf{x}_k^t$  is a positive real scalar, then,

$$\mathbf{x}_k^{tH} (\mathbf{T}_k \boldsymbol{\Phi} \mathbf{T}_k)^{-1} \mathbf{x}_k^t = \operatorname{Re} \left( \operatorname{tr} \left( \mathbf{T}_k^{-1} \boldsymbol{\Phi}^{-1} \mathbf{T}_k^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \right) \right).$$

we obtain,

$$\frac{\partial \operatorname{tr} \left( \mathbf{T}_k^{-1} \boldsymbol{\Phi}^{-1} \mathbf{T}_k^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \right)}{\partial \mathbf{T}_k^{-1}} = 2 \operatorname{Re} \left( \boldsymbol{\Phi}^{-1} \mathbf{T}_k^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \right).$$

Optimizing each  $\hat{\mathbf{T}}_k$  individually results in:

$$\forall k \in [1, M], \hat{\mathbf{T}}_k = \frac{1}{2} \sum_{t=1}^{t=2} \operatorname{Re} \left( \boldsymbol{\Phi}^{-1} \hat{\mathbf{T}}_k^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \right).$$

# Appendix

Let us estimate  $\Phi$  by deriving  $\log(\mathcal{L}_0)$  with respect to  $\Phi^{-1}$ . This involves by deriving the trace of  $(\mathbf{x}_k^{tH} \mathbf{T}_k^{-1}) \Phi^{-1} (\mathbf{T}_k^{-1} \mathbf{x}_k^t)$  with respect to  $\Phi^{-1}$  (see Eq. (101) in [?]):

$$\frac{\partial \log(\mathcal{L}_0)}{\partial \Phi^{-1}} = 2N\Phi - \sum_{\substack{k=1 \\ t=1}}^{k=N \\ t=2} \mathbf{T}_k^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \mathbf{T}_k^{-1} = \mathbf{0}_{mp}.$$

We finally obtain:  $\hat{\Phi} = \frac{1}{2N} \sum_{\substack{k=1 \\ t=1}}^{k=N \\ t=2} \hat{\mathbf{T}}_k^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \hat{\mathbf{T}}_k^{-1}$ .



The same approach is used for  $H_1$ : the texture parameters and the covariance parameter are optimized separately:

$$\log(\mathcal{L}_1) = C - \sum_{\substack{k=1 \\ t=1}}^{\substack{k=N \\ t=2}} \log(|\Phi_t|) + 2 \sum_{\substack{k=1 \\ t=1}}^{\substack{k=N \\ t=2}} \log\left(\left|(\mathbf{T}_k^t)^{-1}\right|\right) - \sum_{\substack{k=1 \\ t=1}}^{\substack{k=N \\ t=2}} \left(\mathbf{x}_k^{tH} (\mathbf{T}_k^t)^{-1} \Phi_t^{-1} (\mathbf{T}_k^t)^{-1} \mathbf{x}_k^t\right).$$

# Appendix

Similarly to  $H_0$ , let us derive  $\log(\mathcal{L}_1)$  with respect to  $(\mathbf{T}_k^t)^{-1}$  and  $\Phi_t^{-1}$ :

$$\frac{\partial \log(\mathcal{L}_1)}{\partial (\mathbf{T}_k^t)^{-1}} = 2 \sum_{\substack{k=1 \\ t=1}}^{k=N} \mathbf{T}_k^t - 2 \sum_{\substack{k=1 \\ t=1}}^{k=N} \operatorname{Re} \left( \Phi_t^{-1} (\mathbf{T}_k^t)^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \right),$$

$$\frac{\partial \log(\mathcal{L}_1)}{\partial \Phi_t^{-1}} = \sum_{\substack{k=1 \\ t=1}}^{k=N} \Phi_t - \sum_{\substack{k=1 \\ t=1}}^{k=N} \mathbf{T}_{tk}^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \mathbf{T}_{tk}^{-1},$$

Letting the two previous equations be equal to  $\mathbf{0}_{m_p}$  leads to the following joint fixed point equations:

$$\begin{cases} \hat{\mathbf{T}}_k^t = \operatorname{Re} \left( \hat{\Phi}_t^{-1} (\hat{\mathbf{T}}_k^t)^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} \right), \\ \hat{\Phi}_t = \frac{1}{N} \sum_{k=1}^{k=N} (\hat{\mathbf{T}}_k^t)^{-1} \mathbf{x}_k^t \mathbf{x}_k^{tH} (\hat{\mathbf{T}}_k^t)^{-1}. \end{cases} \quad (1)$$

# Appendix

## Multiband SAR modeling - Closed-form Solution (two-band case only)

Under  $H_0$ , we have

$$\hat{\mathbf{T}}_k = \begin{pmatrix} \hat{\delta}_{1k} & 0 \\ 0 & \hat{\delta}_{2k} \end{pmatrix},$$

with

$$\begin{cases} \hat{\delta}_{1k}^2 = \frac{1}{2} \sum_{t=1}^{t=2} \left( a_1 + \sqrt{\frac{a_1}{a_2}} a_{12} \right), \\ \hat{\delta}_{2k}^2 = \frac{1}{2} \sum_{t=1}^{t=2} \left( \sqrt{\frac{a_2}{a_1}} a_{12} + a_2 \right). \end{cases}$$

and

$$a_1 = \frac{1}{p} \mathbf{x}_{1,k}^{tH} (\mathbf{M}^{-1})_{11} \mathbf{x}_{1,k}^t, a_2 = \frac{1}{p} \mathbf{x}_{2,k}^{tH} (\mathbf{M}^{-1})_{22} \mathbf{x}_{2,k}^t, a_{12} = \frac{1}{p} \operatorname{Re} \left( \mathbf{x}_{1,k}^{tH} (\mathbf{M}^{-1})_{12} \mathbf{x}_{2,k}^t \right).$$

# Appendix

## Multiband SAR modeling - Closed-form Solution (two-band case only)

Under  $H_1$ , we have:

$$\hat{\mathbf{T}}_{tk} = \begin{pmatrix} \hat{\delta}_{1k}^t & 0 \\ 0 & \hat{\delta}_{2k}^t \end{pmatrix},$$

with

$$\begin{cases} \left(\hat{\delta}_{1k}^t\right)^2 = \hat{\tau}_{1k}^t = a_1 + \sqrt{\frac{a_1}{a_2}} a_{12}, \\ \left(\hat{\delta}_{2k}^t\right)^2 = \hat{\tau}_{2k}^t = a_2 + \sqrt{\frac{a_2}{a_1}} a_{12}, \end{cases}$$

and

$$a_1 = \frac{1}{p} \mathbf{x}_{1,k}^{tH} (\mathbf{M}_t^{-1})_{11} \mathbf{x}_{1,k}^t, a_2 = \frac{1}{p} \mathbf{x}_{2,k}^{tH} (\mathbf{M}_t^{-1})_{22} \mathbf{x}_{2,k}^t, a_{12} = \frac{1}{p} \operatorname{Re} \left( \mathbf{x}_{1,k}^{tH} (\mathbf{M}_t^{-1})_{12} \mathbf{x}_{2,k}^t \right).$$

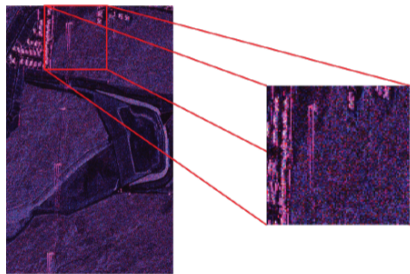
Please note that,  $\forall t \in \{1, 2\}$  and both hypotheses,  $\mathbf{x}^{tH} \hat{\mathbf{T}}_t^{-1} \hat{\Phi}_t^{-1} \hat{\mathbf{T}}_t^{-1} \mathbf{x}^t$  is a positive real scalar:

$$\begin{aligned} \mathbf{x}^{tH} \left( \hat{\mathbf{T}}_t \hat{\Phi}_t \hat{\mathbf{T}}_t \right)^{-1} \mathbf{x}^t &= \operatorname{Re} \left( \operatorname{tr} \left( \left( \hat{\mathbf{T}}_t \hat{\Phi}_t \hat{\mathbf{T}}_t \right)^{-1} \mathbf{x}^t \mathbf{x}^{tH} \right) \right), \\ &= \operatorname{tr} \left( \hat{\mathbf{T}}_t^{-1} \operatorname{Re} \left( \hat{\Phi}_t^{-1} \hat{\mathbf{T}}_t^{-1} \mathbf{x}^t \mathbf{x}^{tH} \right) \right), \\ &= mp. \end{aligned}$$

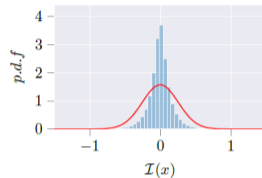
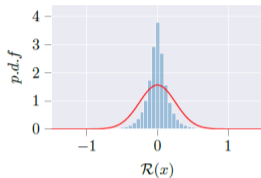
Plugging the estimated parameters leads to the statistic  $\hat{\Lambda}_B(\mathbf{x})$ .

# Appendix

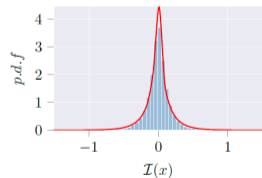
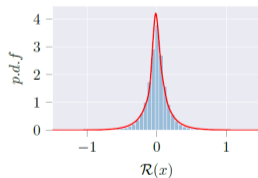
- Illustration of models on real data [1].



High-Resolution SAR image (Left) and a selected subset (Right) for HH polarization.



Gaussian fitting on the selected subset.



Textured Gaussian fitting on the selected subset